

## Distribution Free Confidence Intervals for Quantiles Based on Extreme Order Statistics in a Multi-Sampling Plan

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**Extended Abstract.** Let  $X_{i1}, \dots, X_{in_i}, i = 1, \dots, k$ , be independent random samples from distribution  $F^{\alpha_i}$ , where  $F$  is an absolutely continuous distribution function and  $\alpha_i > 0$ . Also, suppose that these samples are independent. Let  $M_{i,n_i}$  and  $M'_{i,n_i}$ , respectively, denote the maximum and minimum of the  $i$ th sample. Constructing the distribution-free confidence intervals for quantiles of  $F$  based on these informations is the aim of this paper. Various cases have been studied and in each case, the exact non-parametric confidence intervals are obtained. First, we concentrate our attention to the **maxima** of the samples. Coverage probability of a confidence interval based on two different maxima ( $i, j; i \neq j$ ), is derived as follows

$$P(M_{i,n_i} \leq \xi_p \leq M_{j,n_j}) = p^{\alpha_i n_i} (1 - p^{\alpha_j n_j}).$$

For the case of  $P(M_{i,n_i} > M_{j,n_j}) = 1$ , two schemes are suggested as:

1. Using the ordered statistics of  $M_{i,n_i}$  and  $M_{j,n_j}$  as  $M_{ij1} \leq M_{ij2}$ . In this case the associated confidence coefficient is derived as follows

$$P(M_{ij1} \leq \xi_p \leq M_{ij2}) = p^{\alpha_i n_i} (1 - p^{\alpha_j n_j}) + p^{\alpha_j n_j} (1 - p^{\alpha_i n_i}).$$

2. Utilizing the ordered maxima of the  $k$  samples as  $M_{1:k} \leq M_{2:k} \leq \dots \leq M_{k:k}$ . Using  $(M_{i:k}, M_{j:k}), i < j$ , as a confidence interval for  $\xi_p$ , confidence coefficient is obtained, as

$$P(M_{i:k} \leq \xi_p \leq M_{j:k}) = \sum_{r=i}^{j-1} \sum_{A_r} p^{\sum_{s=1}^r n_{t_s} \alpha_{t_s}} \prod_{s=r+1}^k (1 - p^{n_{t_s} \alpha_{t_s}}),$$

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where the summation on  $A_r$  extends over all permutations  $(t_1, \dots, t_k)$  of  $\{1, \dots, k\}$  for which  $t_1 < \dots < t_r$  and  $t_{r+1} < \dots < t_k$ . Several special cases are considered, say, for  $\alpha_i = 1$  and  $n_i = n, i = 1, \dots, k$ , the last identity is simplified as follows

$$P(M_{i:k} \leq \xi_p \leq M_{j:k}) = \sum_{r=i}^{j-1} \binom{k}{r} p^{nr} (1-p)^{k-r},$$

that introduce a binomial representation. By this property, it can be identified  $i$  and  $j$  to achieve the desired qualities. For other special cases such as  $\alpha_i = \frac{1}{n_i}$  or  $\alpha_i = \frac{n^*}{n_i}, i = 1, \dots, k$ , where  $n^* = \sum_{i=1}^k n_i$ , the similar relations are obtained. Similar results for **minima** are obtained. Specially, the following relations are derived

$$P(M'_{ij1} \leq \xi_p \leq M'_{ij2}) = \{1 - (1-p^{\alpha_i})^{n_i}\}(1-p^{\alpha_j})^{n_j} + \{1 - (1-p^{\alpha_j})^{n_j}\}(1-p^{\alpha_i})^{n_i},$$

where  $M'_{ij1} < M'_{ij2}$  denote the ordered minima of two different samples. Also,

$$P(M'_{i:k} \leq \xi_p \leq M'_{j:k}) = \sum_{r=i}^{j-1} \sum_{A_r} \{1 - (1-p^{\alpha_{t_s}})^{n_{t_s}}\} \prod_{s=r+1}^k (1-p^{\alpha_{t_s}})^{n_{t_s}},$$

where  $M'_{1:k} < M'_{2:k} < \dots < M'_{k:k}$  are order statistics of  $M'_{i,n_i}$ 's,  $i = 1, \dots, k$ .

Finally, both **minima** and **maxima** are used to construct the distribution-free confidence intervals for  $\xi_p$  in independent and dependent situations. It is clear that  $M'_{i,n_i}$  and  $M_{j,n_j}, i \neq j$ , are independent and their ordered statistics are denoted by  $V_{ij1} < V_{ij2}$ . Therefore, the following coverage probability is obtained as

$$P(V_{ij1} \leq \xi_p \leq V_{ij2}) = \{1 - (1-p^{\alpha_i})^{n_i}\}(1-p^{\alpha_j n_j}) + p^{\alpha_j n_j} (1-p^{\alpha_i})^{n_i}.$$

In the other hand, by choosing  $(M'_{i,n_i}, M_{i,n_i})$  as a confidence interval, we deal with a dependent end points interval that the confidence coefficient is derived as follows

$$P(M'_{i,n_i} \leq \xi_p \leq M_{i,n_i}) = 1 - p^{\alpha_i n_i} - (1-p^{\alpha_i})^{n_i}.$$

**Keywords.** Extreme order statistics; confidence interval; quantile; coverage probability;  $F^\alpha$  model; reversed hazard rate function.

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